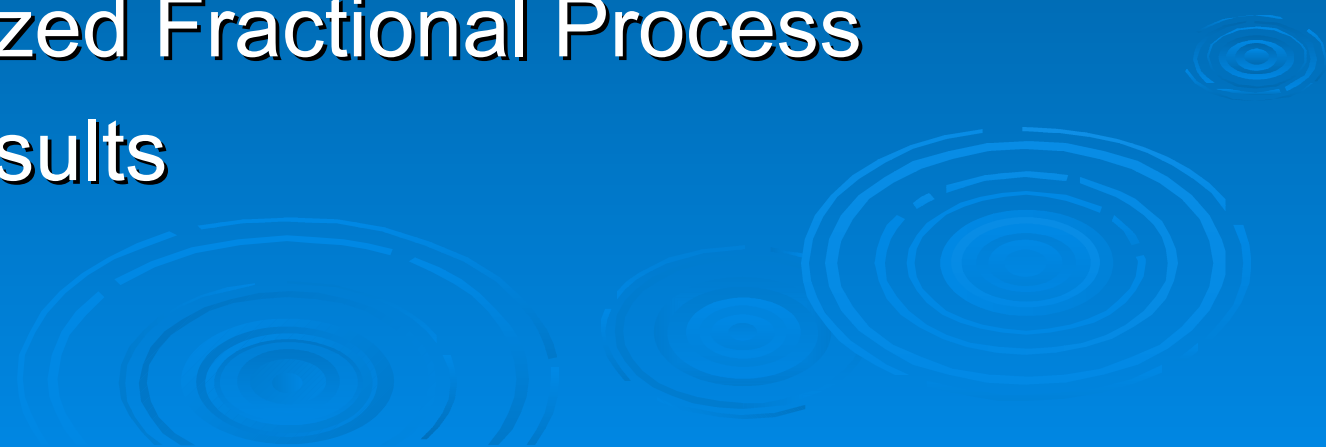


# **Wavelet Analysis of Generalized Fractional Process**

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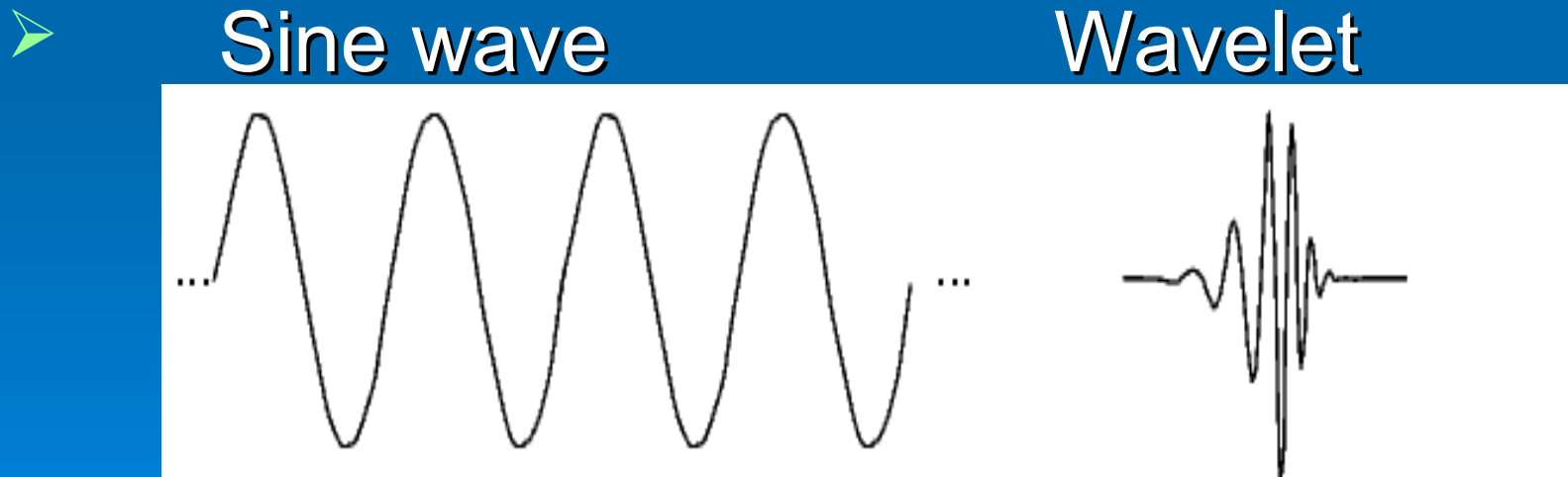
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# What is a wavelet?

- A wavelet is a waveform of effectively limited duration that has an average value of zero.



# What is a wavelet?

## ➤ Admissibility condition:

- The function  $\psi(t) \in L^2(\mathbb{R})$  is often referred to as the *mother wavelet* and must satisfy the admissibility condition given by

- $$\int_{\mathbb{R}} |\Psi(w)|^2 |w|^{-1} dw < \infty,$$

- where  $\Psi(w)$  is the Fourier transform of  $\psi(t)$ .

- If  $\psi(t)$  has sufficient decay, then this condition is equivalent to

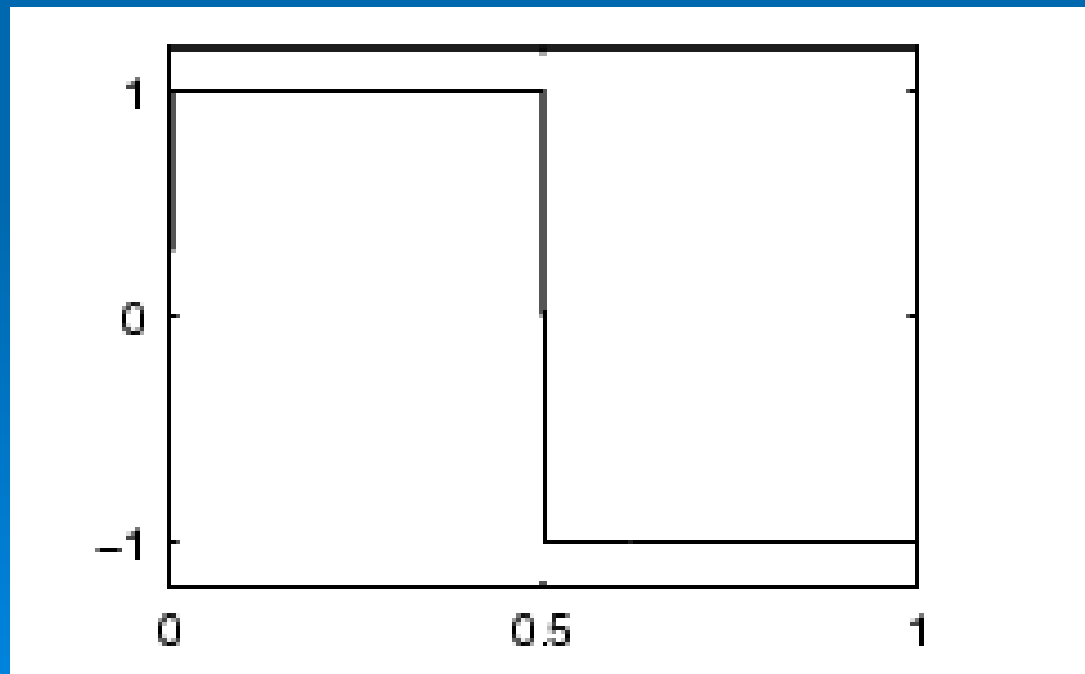
- $$\Psi(0) = \int_{\mathbb{R}} \psi(t) dt = 0.$$



# What is a wavelet?

- Example: Haar wavelet

$$\psi(t) = \begin{cases} 1 & 0 \leq t \leq 1/2 \\ -1 & 1/2 \leq t \leq 1. \\ 0 & \textit{otherwise} \end{cases}$$



# What can wavelets do?

- 1) Performing local analysis
- 2) Analyzing nonstationary signals
- 3) Denoising Signals
- 4) Data Compression
- 5) Decorrelating time series



# Discrete Wavelet Transform

As discretized cwt:

$$C(a, b) = \int_{\mathbb{R}} s(t) \frac{1}{\sqrt{a}} \psi\left(\frac{t-b}{a}\right) dt$$

$$a = 2^j, b = k2^j, (j, k) \in \mathbb{Z}^2$$

# Discrete Wavelet Transform

wavelet coefficients:

$$d_{j,t} = \sum_{l=0}^{L-1} h_l c_{j-1, 2t-1 \bmod N_j-1}$$

scaling coefficients:

$$c_{j,t} = \sum_{l=0}^{L-1} g_l c_{j-1, 2t-1 \bmod N_j-1}$$

$$t = 0, \dots, N_j - 1$$



# Discrete Wavelet Transform

Let  $c_{0,t} = X_t$  (the time series), then the wavelet and scaling coefficients are:

$$d_{j,t} = \sum_{l=0}^{L_j-1} h_{j,l} X_{2^j(t+1)-1-l \bmod N}$$

$$c_{j,t} = \sum_{l=0}^{L_j-1} g_{j,l} X_{2^j(t+1)-1-l \bmod N}$$

$$L_j = (2^j - 1)(L - 1) + 1$$

# Long-memory Process

## Definition:

A *long-memory process* is commonly defined as a stationary process for which the autocorrelation function at lag  $k$  satisfies

$$\rho(k) \sim C_\rho k^{2d-1}$$

as  $k \rightarrow \infty$ , where  $C_\rho \neq 0$  and  $0 < d < 0.5$ .

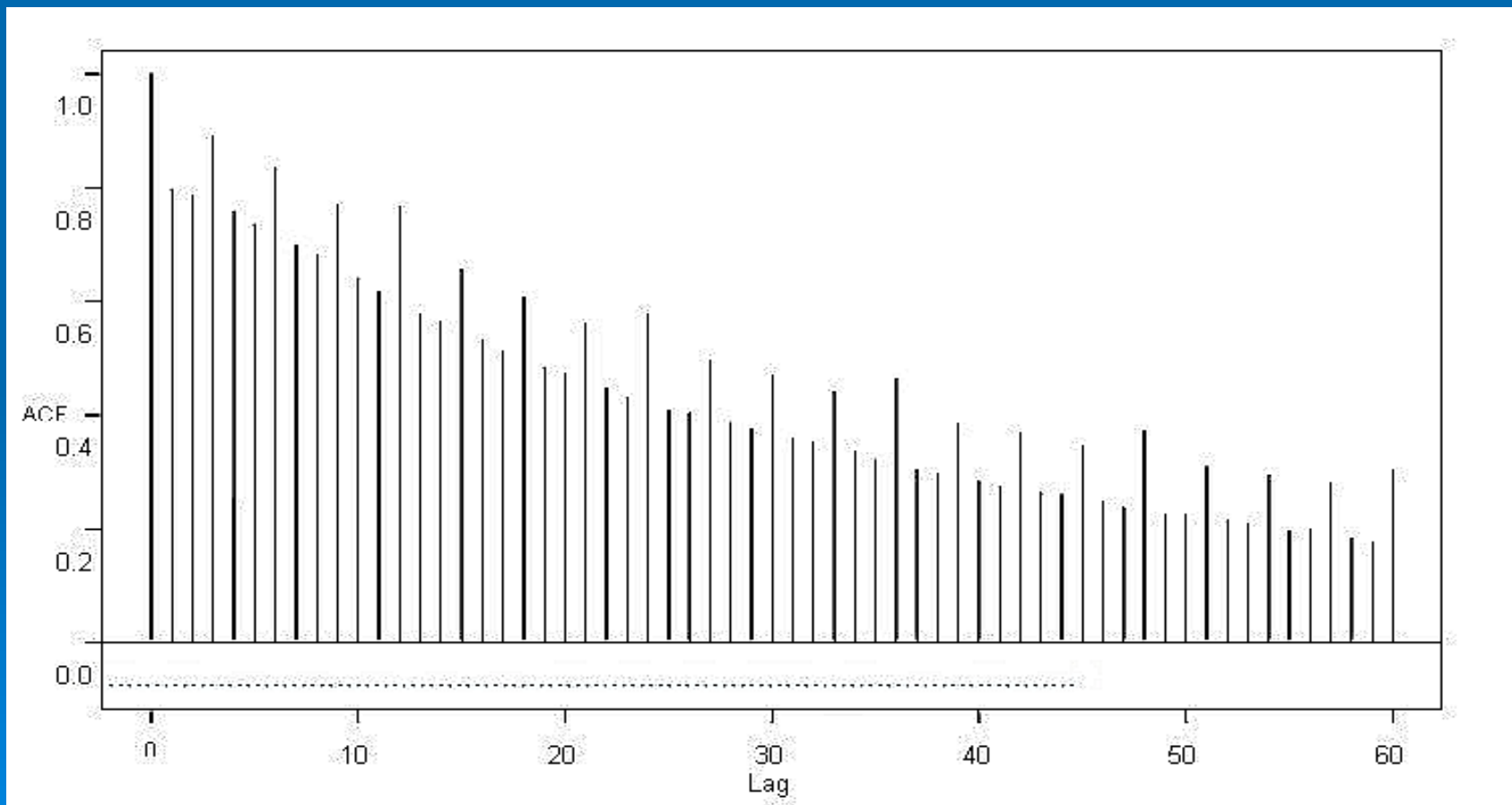
An equivalent statement for the power spectrum is given by

$$f(\omega) \sim C_f |\omega|^{-2d}, \quad \text{as } \omega \rightarrow 0,$$

which has a pole at the origin when  $0 < d < 0.5$ .

# Long-memory Process

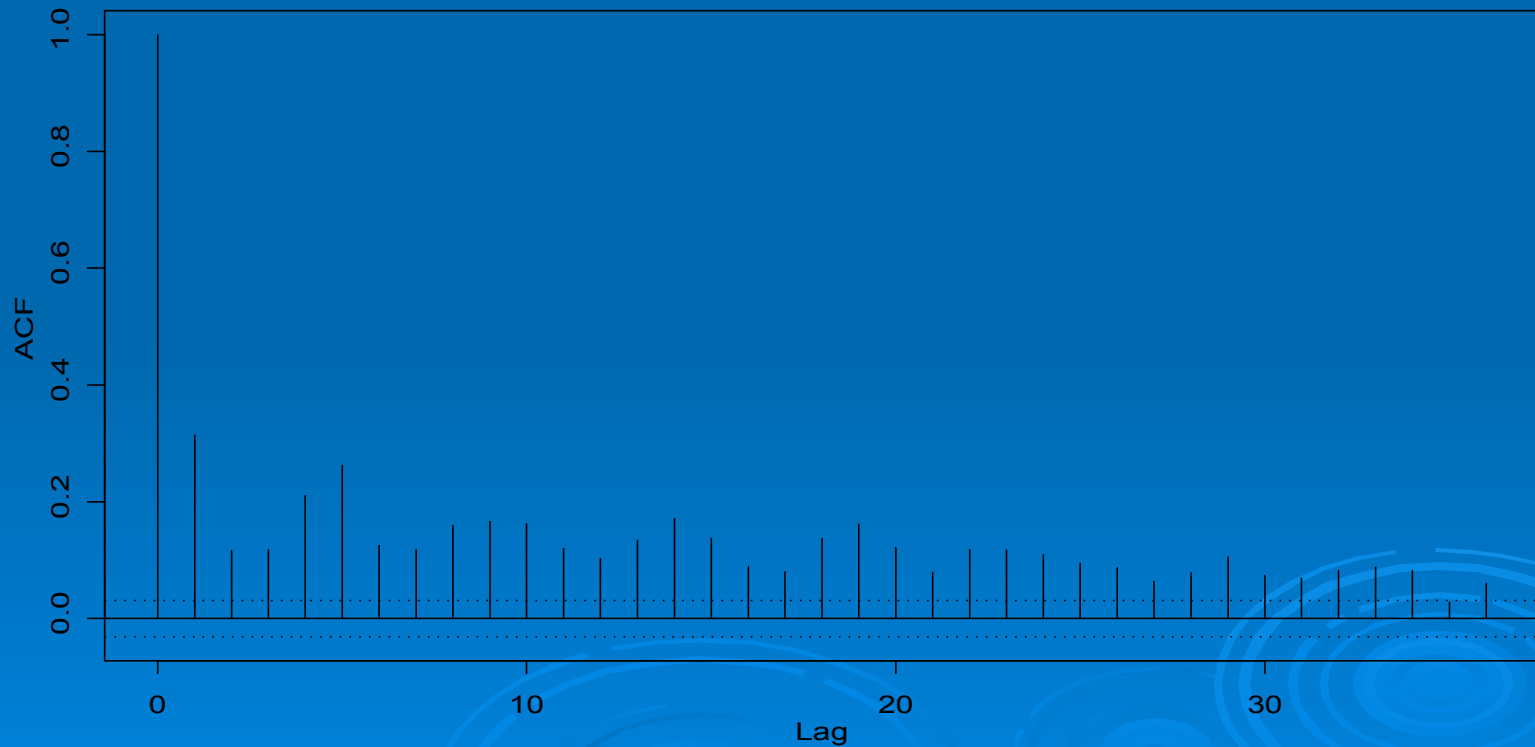
Example: Autocorrelations of video traffic data



# Long-memory Process

Example: Autocorrelations of Ethernet data

Series : Ethernet



# Generalized Fractional Process

## The Model

A *Gegenbauer autoregressive moving average*  $GARMA(p,d,u,q)$  process is the output of the system function

$$H(z) = \frac{\Theta(z)}{\Phi(z)} (1 - 2uz^{-1} + z^{-2})^{-d}$$

driven by a stationary white noise input with mean 0 and variance  $\sigma^2$ .

# Generalized Fractional Process

## The Model

The rational function

$$\frac{\Theta(z)}{\Phi(z)} = \frac{1 + \theta_1 z^{-1} + \dots + \theta_q z^{-q}}{1 - \phi_1 z^{-1} - \dots - \phi_p z^{-p}}$$

is the *autoregressive moving average*,  $ARMA(p,q)$  system, such that  $z^p \Theta(z)$  and  $z^p \Phi(z)$  have no common zeros and the zeros lie outside the unit circle.

# Generalized Fractional Process

## The Model

The *Gegenbauer* system is defined by

$$\left(1 - 2uz^{-1} + z^{-2}\right)^{-d} = \sum_{n=0}^{\infty} C_n^d(u) z^{-n}$$

$$C_n^d(u) = \sum_{k=0}^{[n/2]} \frac{(-1)^k (2u)^{n-2k} \Gamma((d) - k + n)}{k!(n-2k)!\Gamma(d)}$$

When the input is a stationary white noise, the output is called a *Gegenbauer process*, which is stationary if  $d < 0.5$  and  $|u| < 1$  or if  $d < 0.25$  and  $|u| = 1$ ; it is invertible if  $-0.5 < d$  and  $|u| < 1$  or  $-0.25 < d$  and  $|u| = 1$ . If  $u = 1$ , we have ARFIMA(p,d,q) process.

# Generalized Fractional Process

## The Model

The power spectrum of a GARMA( $p, d, u, q$ ) process is given by

$$f(\omega) = \sigma_z^2 \left| \frac{\Theta(e^{-i2\pi\omega})}{\Phi(e^{-i2\pi\omega})} \right|^2 \left[ 4(\cos(2\pi\omega) - u)^2 \right]^{-d}$$

where  $\omega \in (-0.5, 0.5]$  and  $\nu = \cos^{-1}(u)/2\pi$  is called the *Gegenbauer frequency* at which the power spectrum becomes unbounded when  $0 < d < 0.5$ .



# Generalized Fractional Process

## The Model

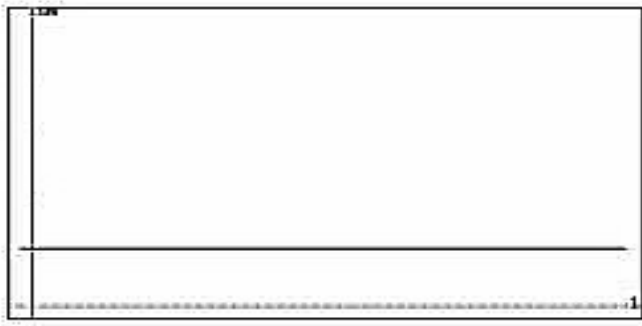
The autocovariance function of a GARMA( $p, d, u, q$ ) process is given by

$$\gamma(k) = \frac{\sigma^2}{2\sqrt{\pi}} \Gamma(1-2d) [2 \sin(v)]^{0.5-2d} \cdot \left[ P_{k-0.5}^{2d-0.5}(u) + (-1)^k P_{k-0.5}^{2d-0.5}(-u) \right]$$

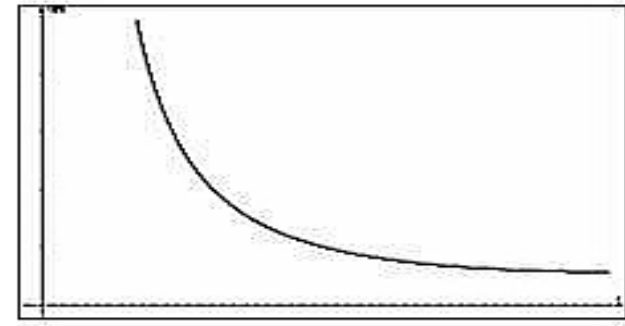
where  $P_k^m$  are the *associated Legendre functions of the first kind*.

# Generalized Fractional Process

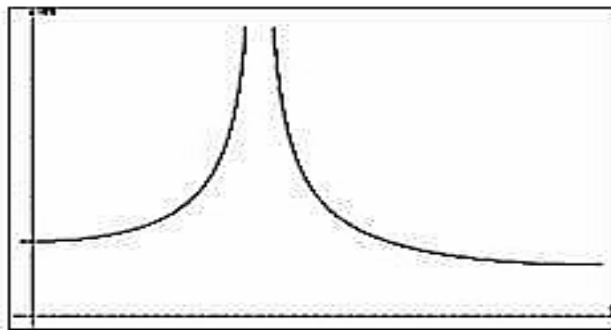
Spectrum of Gamma(0,d,u,0) Process



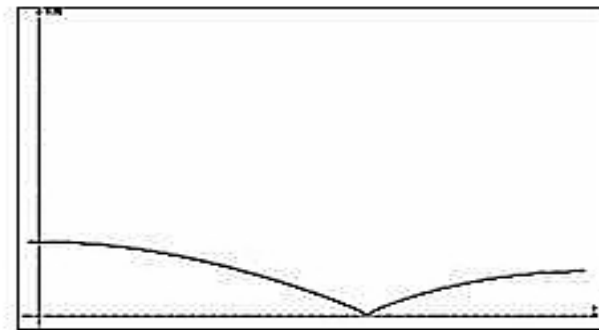
(a)



(b)



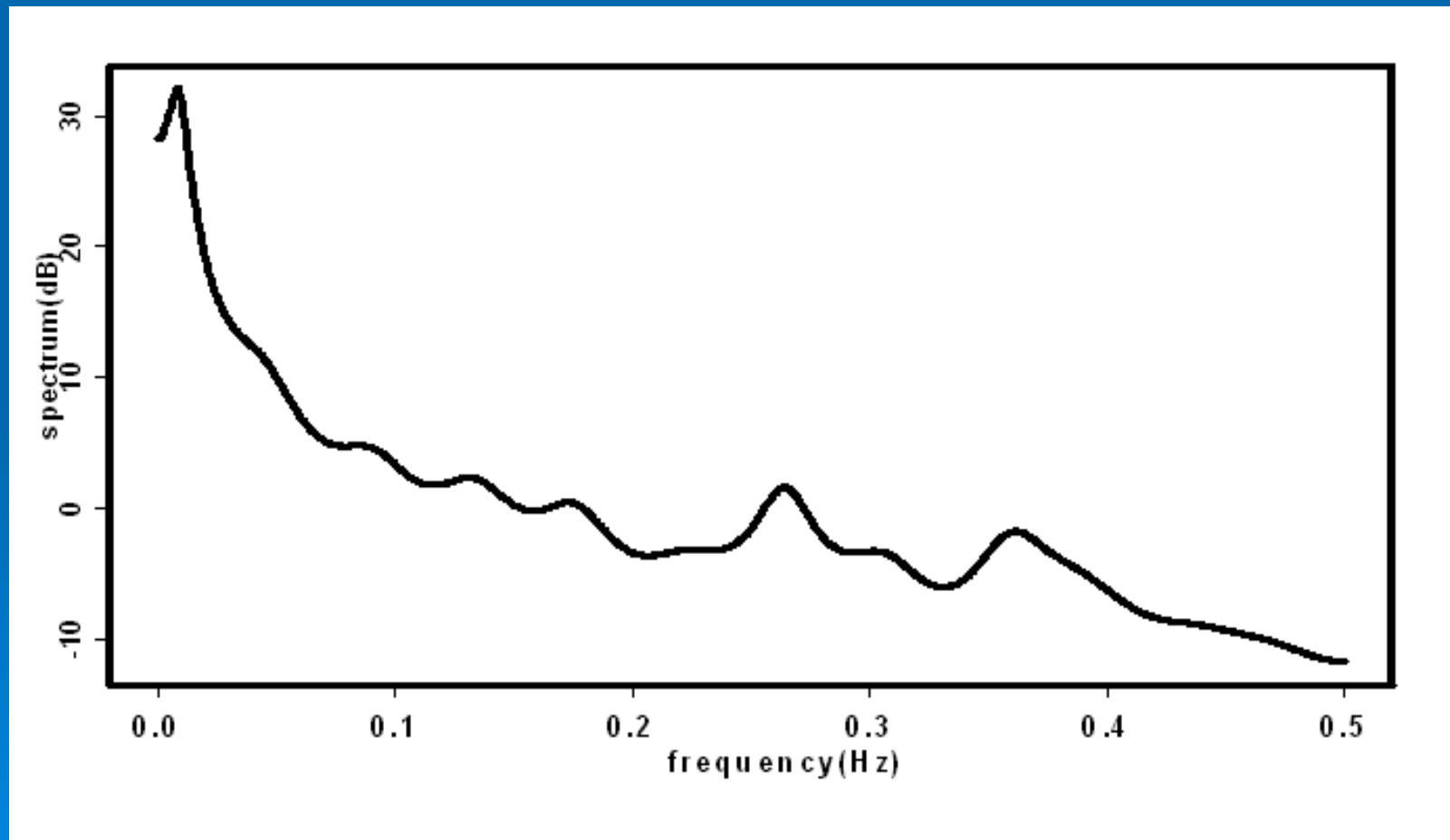
(c)



(d)

# Generalized Fractional Process

Example: Spectrum of heart rate data



# Generalized Fractional Process

GARMA(p,d,u,q) Process generalizes the following:

- 1) Gegenbauer process
- 2) Fractionally Integrated process
- 3) ARMA Process
- 4) AR Process
- 5) MA Process

# Covariance Structure of Wavelet Coefficients

## Covariance of wavelet coefficients

$$\text{Cov}(d_{jt}, d_{j't'}) = \int_{-1/2}^{1/2} e^{i2\pi f(2^{j'}(t'+1) - 2^j(t+1))} H_j(f) H_{j'}^*(f) S_Y(f) df$$

$$\text{Cov}(d_{jt}, d_{j(t+s)}) = \int_{-1/2}^{1/2} e^{i2\pi fs} |H_j(f)|^2 S_Y(f) df$$

where  $H_{j,L}(f)$  is the Fourier transform of the level  $j$  Daubechies wavelet filter  $\{h_{j,l}\}$

# Covariance Structure of Wavelet Coefficients

Note that:

$$H_{j,L}(f) = H_{1,L}(2^{j-1}f) \prod_{l=0}^{j-2} G_{1,L}(2^l f)$$

$$G_{j,L}(f) = \prod_{l=0}^{j-1} G_{1,L}(2^l f)$$

# Covariance Structure of Wavelet Coefficients

**Lemma 1** (Gonzaga and Kawanaka) Let  $\{h_{\ell,1}, \ell = 0, 1, \dots, L-1\}$  the orthonormal Daubechies wavelet filter of length  $L$ , then as  $L \rightarrow \infty$ ,

$$|H_{j',L}(f)|^2 |G_{1,L}(2^{j'-1}f)|^2 \rightarrow 0 \text{ a.e. and}$$

$$|H_{1,L}(2^{j'-1}f)|^2 \prod_{m=0}^{j-2} |G_{1,L}(2^m f)|^2 \rightarrow 0 \text{ a.e.} \quad (\text{A1})$$

on  $[0, 0.5]$

# Covariance Structure of Wavelet Coefficients

Theorem 2 (Gonzaga and Kawanaka). Let  $\{Y_t\}$  be a generalized fractional process and  $\{h_{\ell,1}, \ell = 0, 1, \dots, L-1\}$  the orthonormal Daubechies wavelet filter of length  $L$ , then for  $j > j'$  and  $d < 0$ ,

$$|\text{Cov}(d_{j,t}, d_{j',t'})| = O\left(\frac{1}{L^{3/4}}\right). \quad (\text{A7})$$



# Covariance Structure of Wavelet Coefficients

Theorem 3 (Gonzaga and Kawanaka). Let  $\{Y_t\}$  be a generalized fractional process and  $\{h_{\ell,1}, \ell = 0, 1, \dots, L-1\}$  the orthonormal Daubechies wavelet filter of length  $L$ , then for  $j > j'$  and  $d > 0$ ,

$$|\text{Cov}(d_{jt}, d_{j't'})| = O\left(\frac{1}{L^{3/4}}\right), \quad \text{if } v \in [0, 2^{-j-1}], \quad (\text{A24})$$

and

$$|\text{Cov}(d_{jt}, d_{j't'})| = O\left(\frac{1}{L^{1/4}}\right), \quad \text{if } v \notin [0, 2^{-j-1}]. \quad (\text{A25})$$

# Covariance Structure of Wavelet Coefficients

**Theorem 4** (Gonzaga and Kawanaka) Let  $\{Y_t\}$  be a generalized fractional process and  $\{h_{\ell,1}, \ell = 0, 1, \dots, L-1\}$  the orthonormal Daubechies wavelet filter of length  $L$ , then

$$\lim_{L \rightarrow \infty} \text{cov}(d_{jt}, d_{j(t+s)}) = 2^{j+1} \int_{2^{-j-1}}^{2^{-j}} \cos(2^{j+1} \pi f s) S_Y(f) df, \quad (\text{A52})$$

which exists for  $j \in Z^+$ .

# Covariance Structure of Wavelet Coefficients

Theorem 5 (Gonzaga and Kawanaka). Let  $\{Y_t\}$  be a generalized fractional process and  $\{h_{\ell,1}, \ell = 0, 1, \dots, L-1\}$  the orthonormal Daubechies wavelet filter of length  $L$ , then if  $u=-1$  and  $j>1$

$$|\text{cov}(d_{jt}, d_{j(t+s)})| = O([2^j s]^{-[L-4d]-1}) \text{ as } 2^j s \rightarrow \infty. \quad (\text{A59})$$

# Covariance Structure of Wavelet Coefficients

**Lemma 6.** (Gonzaga and Kawanaka) Let  $\{h_{\ell,1}, \ell = 0, 1, \dots, L-1\}$  be the orthonormal Daubechies wavelet filter of length  $L$  and  $|H_{j,L}(f)|^2$  its energy spectrum at level  $j$  and frequency  $f$ . Then

$$|H_{j,L}(f)|^2 \leq \frac{1}{2^j} \left( \frac{2 \sin^2(2^{j-1} \pi f)}{\sin(\pi f)} \right)^{L-2}. \quad (\text{A66})$$

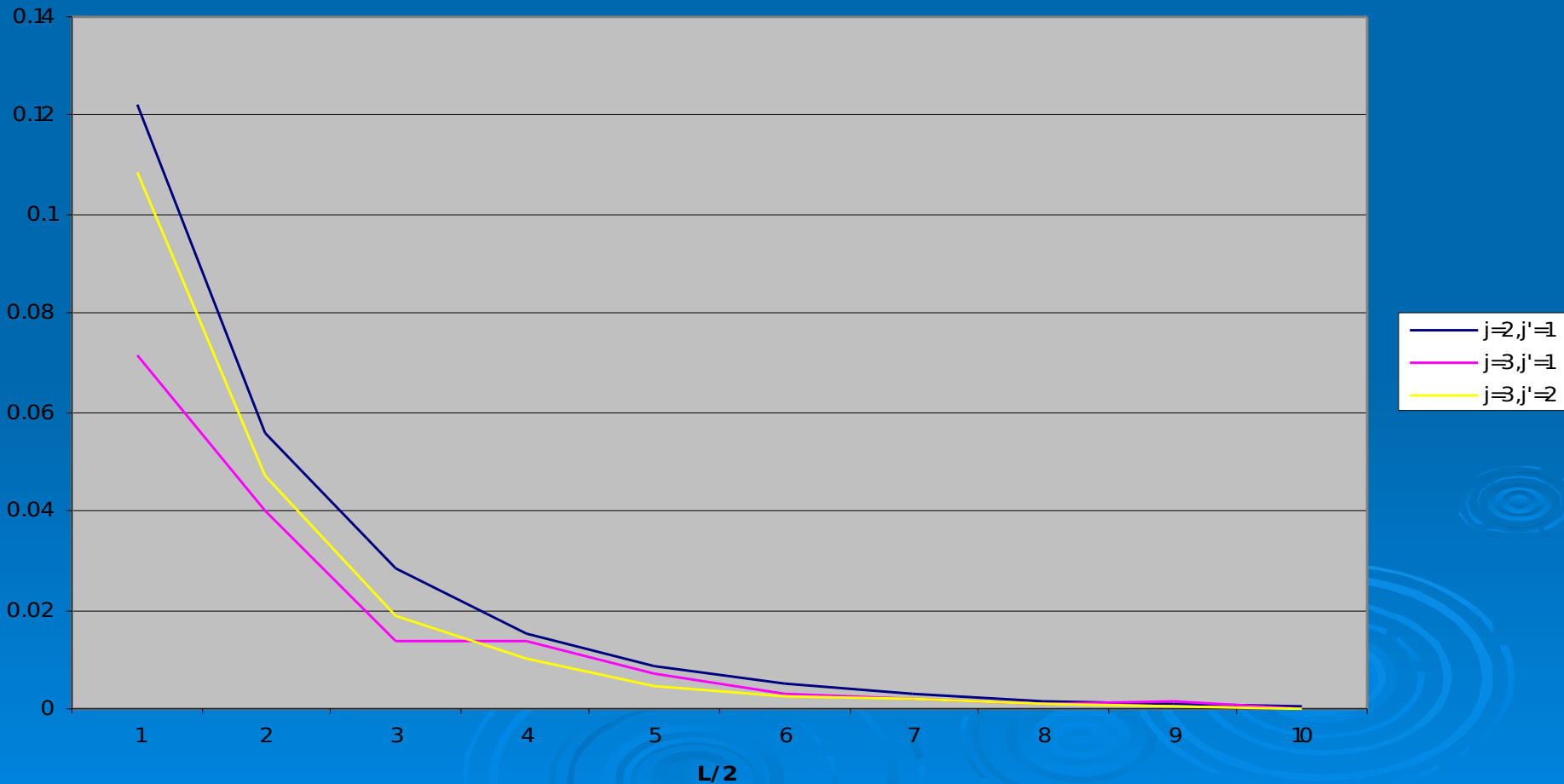
# Covariance Structure of Wavelet Coefficients

Theorem 7 (Gonzaga and Kawanaka). Let  $\{Y_t\}$  be a generalized fractional process and  $\{h_{\ell,1}, \ell = 0, 1, \dots, L-1\}$  the orthonormal Daubechies wavelet filter of length  $L$ , then if  $u \in (-1, 1)$

$$|\text{Cov}(d_{j,t}, d_{j,t+s})| = O\left(2^{j(2d-2)} s^{2d-1}\right) \text{ as } 2^j \rightarrow \infty \text{ and } s \rightarrow \infty. \quad (\text{A78})$$

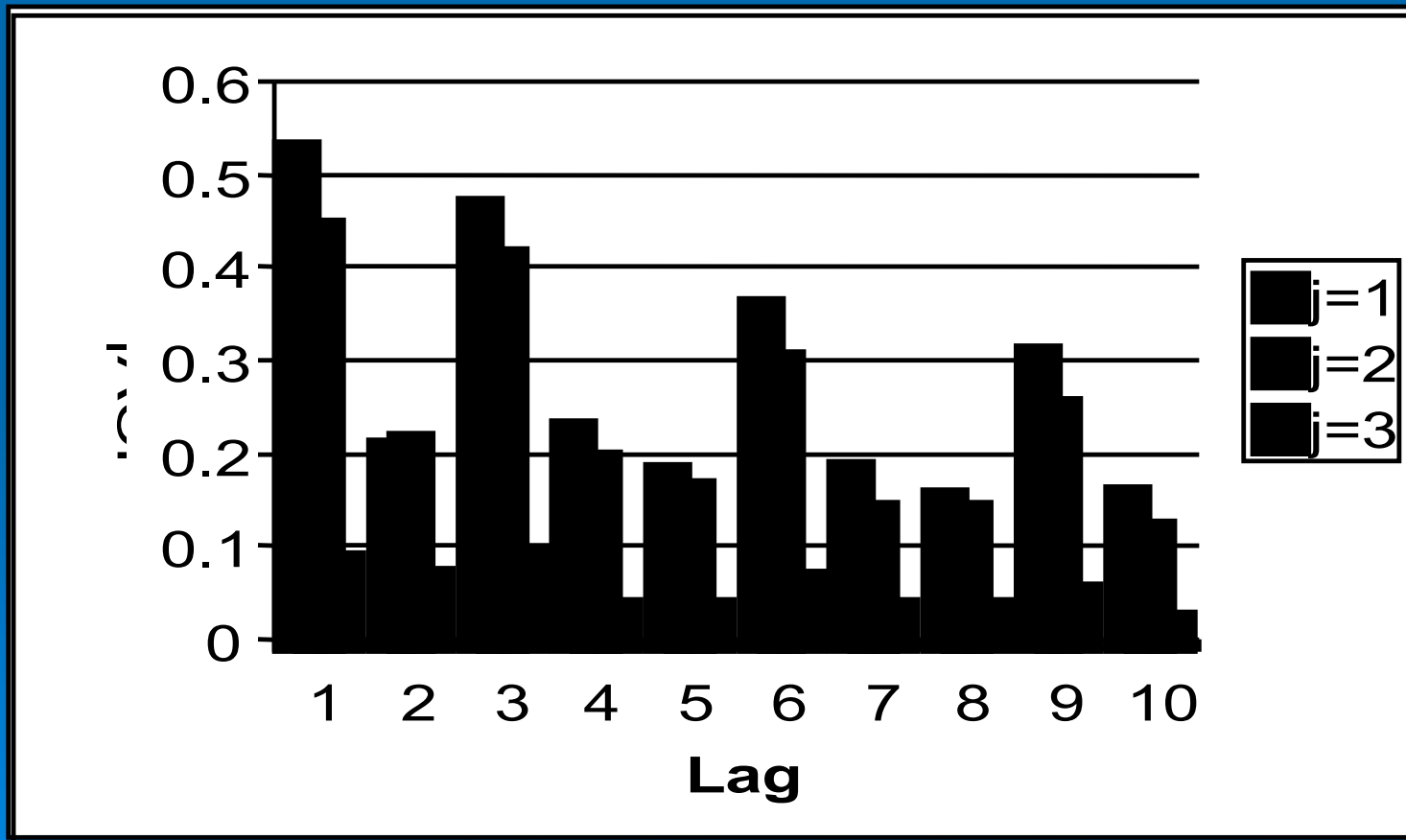
# Covariance Structure of Wavelet Coefficients

Absolute maximum values of correlations



# Covariance Structure of Wavelet Coefficients

Within-scale correlations



# Weighted Least Square Estimation

Wavelet variance

$$v_Y^2(\lambda) = \frac{E(W_{t,\lambda}^2)}{2\lambda}$$

Maximal overlap estimator

$$\hat{v}_Y^2(\lambda) = \frac{1}{2\lambda N_{W_\lambda}} \sum_{t=L_\lambda}^N w_{t,\lambda}^2$$

Note:

$$\log \hat{v}^2(\lambda) \xrightarrow{d} N\left(\log v^2(\lambda), A_{W_\lambda} / (2\lambda^2 N_{W_\lambda} v^4(\lambda))\right)$$



# Weighted Least Square Estimation

Regression equation:

$$\log(v^2(\lambda_j)) \approx -2d \log(2|\cos(2\pi \mu) - \cos(2\pi v)|)$$

# Weighted Least Square Estimation

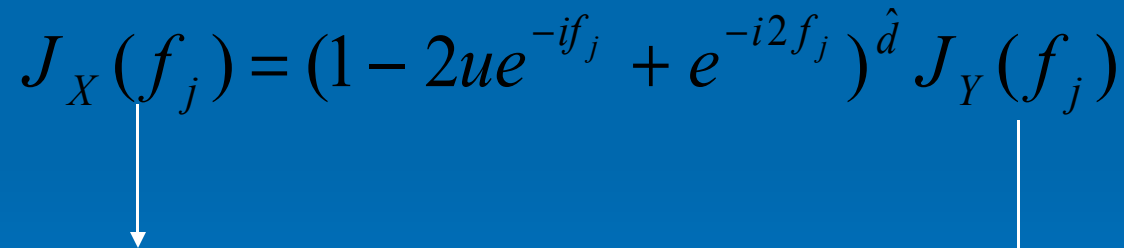
Estimator of the long-memory parameter

$$\hat{d} = -\frac{1}{2} \left[ \frac{\sum_{j=1}^J u_j x_j y_j - \left( \sum_{j=1}^J u_j y_j \right) \left( \sum_{j=1}^J u_j x_j \right)}{\sum_{j=1}^J u_j x_j^2 - \left( \sum_{j=1}^J u_j x_j \right)^2} \right]$$

$$x_j = \log\left(2\left|\cos(2\pi\mu_j) - \cos(2\pi\nu)\right|\right) \quad y_j = \log\left(\hat{v}^2(2^j)\right)$$

# Weighted Least Square Estimation

Estimation of short-memory parameters:

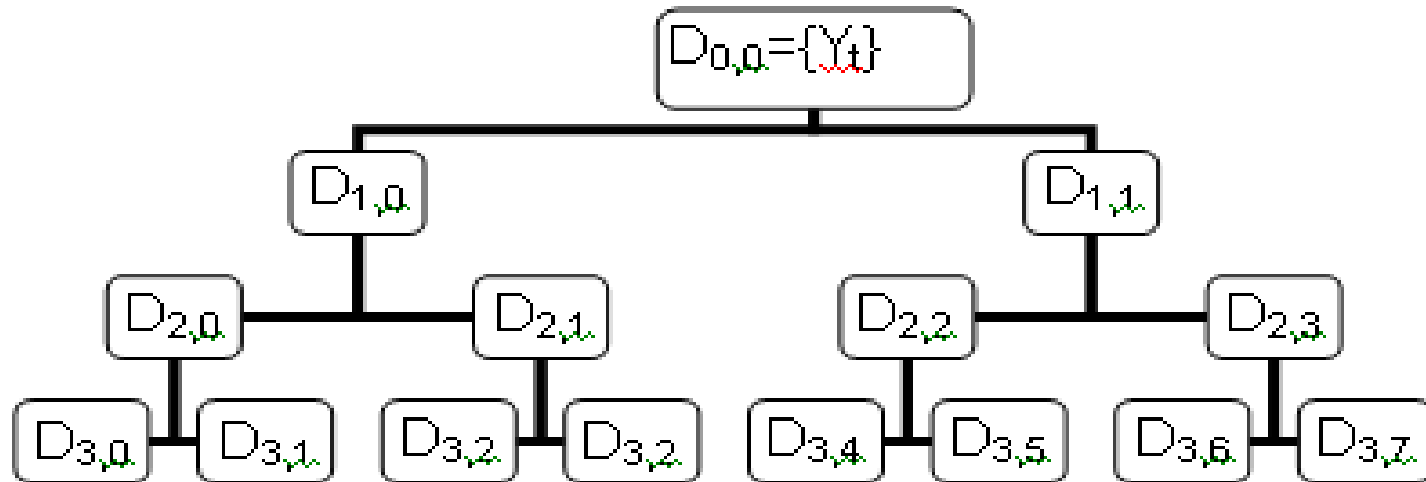
$$J_X(f_j) = (1 - 2ue^{-if_j} + e^{-i2f_j})^{\hat{d}} J_Y(f_j)$$


DFT of ARMA(p,q)

DFT of GARMA(p,d,u,q)

# Likelihood Estimation

- We use wavelet packet wavelet transform



Wavelet packet decomposition

# Likelihood Estimation

## Basis Selection Algorithm

For  $j < J$ , we test the vector  $D_{j,n}$  for white noise. If the test fails to reject, we retain  $D_{j,n}$ . If the test rejects, we split  $D_{j,n}$  into  $D_{j+1,2n}$  and  $D_{j+1,2n+1}$ , and test both the resulting subbands for white noise. We repeat this process until  $j=J$  in which we retain  $D_{J,n}$ . We denote the resulting vector of DWPT coefficients by  $\mathbf{D} = (D_{j,n}, (j, n) \in B)$ , which is approximately uncorrelated.

# Likelihood Estimation

The approximate likelihood can be written as a univariate density from which e.g. the posterior density can be obtained and an MCMC algorithm be implemented:

$$L(D | \Psi) = \left(2\pi\sigma_{\varepsilon}^2\right)^{-N/2} \left( \prod_{(j,n) \in B} \left(\sigma_{jn}^2\right)^{-N_{jn}/2} \right) \exp \left[ \frac{-1}{2\sigma_{\varepsilon}^2} \sum_{(j,n) \in B} \frac{D_{j,n}^T D_{j,n}}{\sigma_{jn}^2} \right]$$

End

