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Monitoring Structural Change in Dynamic Econometric Models

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Model frame

- Generalized fluctuation tests
 - ♦ OLS-based processes
 - Rescaling of estimates-based processes
 - Boundaries
- Applications
 - German M1 money demand
 - ♦ U.S. labor productivity



Consider the linear regression model in a monitoring situation

$$y_i = x_i^\top \beta_i + u_i$$
 $(i = 1, \dots, n, \dots),$

where at time *i*:

* y_i — dependent variable,

 x_i — vector of k regressors,

 $\ \ \beta_i - vector of k unknown regression coefficients,$

 $* u_i$ — error term.



It is assumed that the regression relationship is stable ($\beta_i = \beta_0$) during the history period i = 1, ..., n.

Null hypothesis:

$$H_0: \quad \beta_i = \beta_0 \qquad (i > n),$$

Alternative:

$$H_1: \quad \beta_i \neq \beta_0 \qquad \text{for some } (i > n).$$

Generalized fluctuation tests **TU**

- * empirical fluctuation processes reflect fluctuation in
 - ✤ residuals
 - coefficient estimates
- theoretical limiting process is known
- * choose boundaries which are crossed by the limiting process only with a known probability α .
- * if the empirical fluctuation process crosses the theoretical boundaries the fluctuation is improbably large \Rightarrow reject the null hypothesis.

Generalized fluctuation tests **TU**

Chu, Stinchcombe, White (1996)

Extension of fluctuation tests to the monitoring situation: processes based on recursive estimates and recursive residuals.

Leisch, Hornik, Kuan (2000)

Generalized framework for estimates-based tests for monitoring.

Contains the test of Chu et al., and considered in particular moving estimates.

Processes based on estimates:

$$\hat{\beta}^{(i)} = \left(X_{(i)}^{\top} X_{(i)}\right)^{-1} X_{(i)}^{\top} y^{(i)}$$

Recursive estimates (RE) process:

$$Y_n(t) = \frac{i}{\widehat{\sigma}\sqrt{n}} Q_{(n)}^{\frac{1}{2}} \left(\widehat{\beta}^{(i)} - \widehat{\beta}^{(n)}\right),$$

where $i = \lfloor k + t(n-k) \rfloor$ and $t \ge 0$.

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Moving estimates (ME) process:

$$Z_n(t|h) = \frac{\lfloor nh \rfloor}{\widehat{\sigma}\sqrt{n}} Q_{(n)}^{\frac{1}{2}} \left(\widehat{\beta}^{(\lfloor nt \rfloor - \lfloor nh \rfloor, \lfloor nh \rfloor)} - \widehat{\beta}^{(n)}\right),$$

where $t \geq h$.

The empirical processes converge to a k-dimensional Brownian bridge or the increments thereof respectively.

The null hypothesis is rejected when the empirical processes cross the boundary

$$b_1(t) = \sqrt{t(t-1)\left[\lambda^2 + \log\left(\frac{t}{t-1}\right)\right]}$$

or

$$c(t) = \lambda \cdot \sqrt{\log_+ t}$$

respectively in the monitoring period 1 < t < T and λ determines the significance level of this procedure.



Processes based on OLS residuals:

$$\hat{u}_i = y_i - x_i^\top \hat{\beta}^{(n)}$$

OLS-based CUSUM process:

$$W_n^0(t) = \frac{1}{\hat{\sigma}\sqrt{n}} \sum_{i=1}^{\lfloor nt \rfloor} \hat{u}_i \qquad (t \ge 0).$$



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$$W_n^0(t) = \frac{1}{\widehat{\sigma}\sqrt{n}} \sum_{i=1}^{\lfloor nt \rfloor} \widehat{u}_i \qquad (t \ge 0).$$

OLS-based MOSUM process:

$$M_n^{0}(t|h) = \frac{1}{\widehat{\sigma}\sqrt{n}} \left(\sum_{i=\lfloor \eta t \rfloor - \lfloor nh \rfloor + 1}^{\lfloor \eta t \rfloor} \widehat{u}_i \right) \qquad (t \ge h).$$



The limiting processes are the 1-dimensional Brownian bridge or the increments thereof respectively. Thus, the same boundaries can be used.

Advantage: ease of computation.



Kuan & Chen (1994):

Empirical size of (historical) estimates-based tests can be seriously distorted in dynamic models if the whole sample covariance matrix estimate

$$Q_{(n)} = 1/n \cdot X_{(n)}^{\top} X_{(n)}$$

is used to scale the fluctuation process. Improvement: use $Q_{(i)}$ instead.

In a monitoring situation rescaling cannot improve the size of the RE test but it does so for the ME test!



Example: AR(1) process with $\rho = 0.9$ but *without* a shift:





The shape of the boundaries does not make a big difference under H_0 , but determines the power for certain alternatives.

Standard boundaries:





Consider a boundary with an offset in t = 1, but with the correct asymptotical growth rate $t \Rightarrow$ the simplest case:

$$b_2(t) = \lambda \cdot t.$$





This spreads the size much more evenly:



Lütkepohl, Teräsvirta, Wolters (1999) investigate the linearity and stability of German M1 money demand: stable regression relation for the time before the monetary union on 1990-06-01 but a clear structural instability afterwards.

Data: seasonally unadjusted quarterly data, 1961(1) to 1995(4)

Error Correction Model (in logs) with variables: M1 (real, per capita) m_t , price index p_t , GNP (real, per capita) y_t and long-run interest rate R_t :

$$\Delta m_t = -0.30 \Delta y_{t-2} - 0.67 \Delta R_t - 1.00 \Delta R_{t-1} - 0.53 \Delta p_t$$

-0.12m_{t-1} + 0.13y_{t-1} - 0.62R_{t-1}
-0.05 - 0.13Q1 - 0.016Q2 - 0.11Q3 + \hat{u}_t ,

Historical OLS-based tests...do *not* discover shift:



The shift has an estimated angle of 90.27° .

Historical estimates-based tests discover shift *ex post*:



Time

Monitoring with OLS-based CUSUM test

Monitoring discovers shift *online*:



Monitoring discovers shift *online*:



Monitoring with OLS-based CUSUM test

Monitoring with OLS-based CUSUM test

Monitoring discovers shift *online*:





Hansen (2001) examines U.S. labor productivity in the manufacturing/durables sector

Monthly data, 1947(2) to 2001(4), AR(1) model.

Finds a clear structural change in about 1994 and two weaker changes in 1963 and 1982.

History period: 1964(1) to 1979(12)

$$x_t = 0.0025 - 0.186x_{t-1} + \hat{u}_t$$



Monitoring with OLS-based CUSUM test



Time



Monitoring with OLS-based CUSUM test



Time



Monitoring with OLS-based CUSUM test



Time



= 1000 blocess

Monitoring with ME test (moving estimates test)

U.S. labor productivity



Monitoring with ME test (moving estimates test)





All methods implemented in R

http://www.R-project.org/

in the contributed package strucchange available from the Comprehensive R Archive Network (CRAN):

http://cran.R-project.org/

documented in:

A. Zeileis, F. Leisch, K. Hornik, C. Kleiber (2002), "strucchange: An R Package for Testing for Structural Change in Linear Regression Models," *Journal of Statistical Software*, 7(2), 1–38.